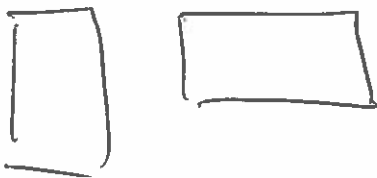
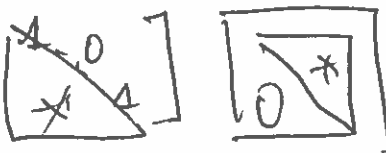



# Singular value decomposition of $A \in M_{m,n}$ .

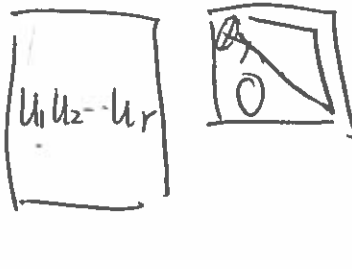
I  $A = CR =$  

$A = \begin{bmatrix} 1 & 1 \\ 1 & 1+\delta \end{bmatrix}$

$Ax = b$   
 $x = A^{-1}b$

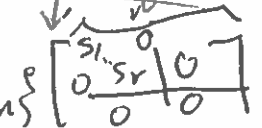
II  $A = LU =$  

III  $PA = LU =$  

IV  $A = QR =$  

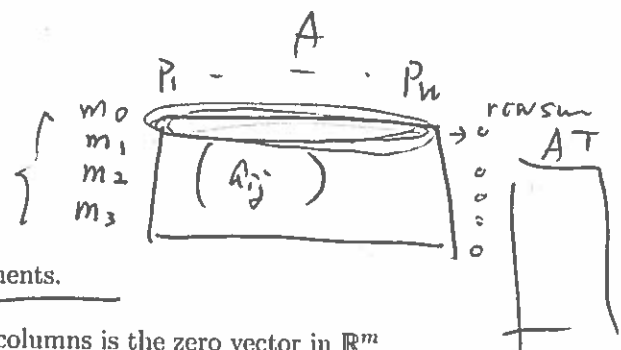
$n \times r$   
 $\begin{bmatrix} | & | & | \\ A_1 & A_2 & \dots \\ | & | & | \end{bmatrix}$

V  $S^{-1}AS = \begin{bmatrix} \lambda & & \\ & \lambda & \\ & & \lambda \end{bmatrix}$   $A = SDS^{-1}$   $D = \begin{bmatrix} d_1 & & 0 \\ & \ddots & \\ 0 & & d_n \end{bmatrix}$   
 $n$  possible when there are  $n$  linearly independent eigenvectors.  
 $M_n$

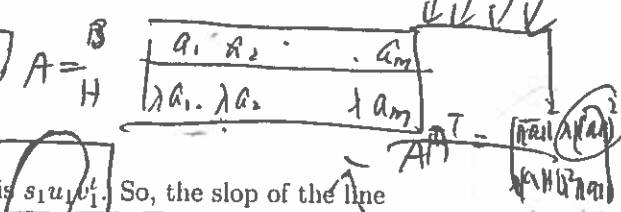
VI SVD of  $A \in M_{m,n}$ .  $A = U \Sigma V^*$   
 $U$  is  $m \times m$  unitary,  $\Sigma$  is  $m \times n$ ,  $V^*$  is  $n \times n$ .  
  
 $s_1 \geq \dots \geq s_r > 0$   
 $= s_1 u_1 v_1^* + \dots + s_r u_r v_r^*$

Best rank  $k$  approximation of  $A$  is  
 $= s_1 u_1 v_1^* + \dots + s_k u_k v_k^*$

# Principal Component Analysis



- Let  $A$  be  $m \times n$ , each column is a sample with  $m$  measurements.
- Normalize the means to 0 for all measurements. So, sum of columns is the zero vector in  $\mathbb{R}^m$ .
- The variances are the diagonal entries of  $AA^T \in M_m$ .
- The co-variances are the off-diagonal entries of  $AA^T \in M_m$ .
- The sample co-variance matrix is  $S = AA^T / (n - 1)$ .



For example, when  $m = 2$ , the best rank one approximation of  $S$  is  $s_1 u_1 u_1^T$ . So, the slope of the line is the ratio of the second entry of  $u_1$  to the first entry of  $u_1$ .

If the rank of  $AA^T$  is low, then the two measurement is closely related, i.e., almost agree on a linear relation.

We can extend the idea to higher dimension data set recorded as  $A \in M_{m,n}$ . One can use  $k$ -dimensional hyperplane to approximate the data.

- The total variance is

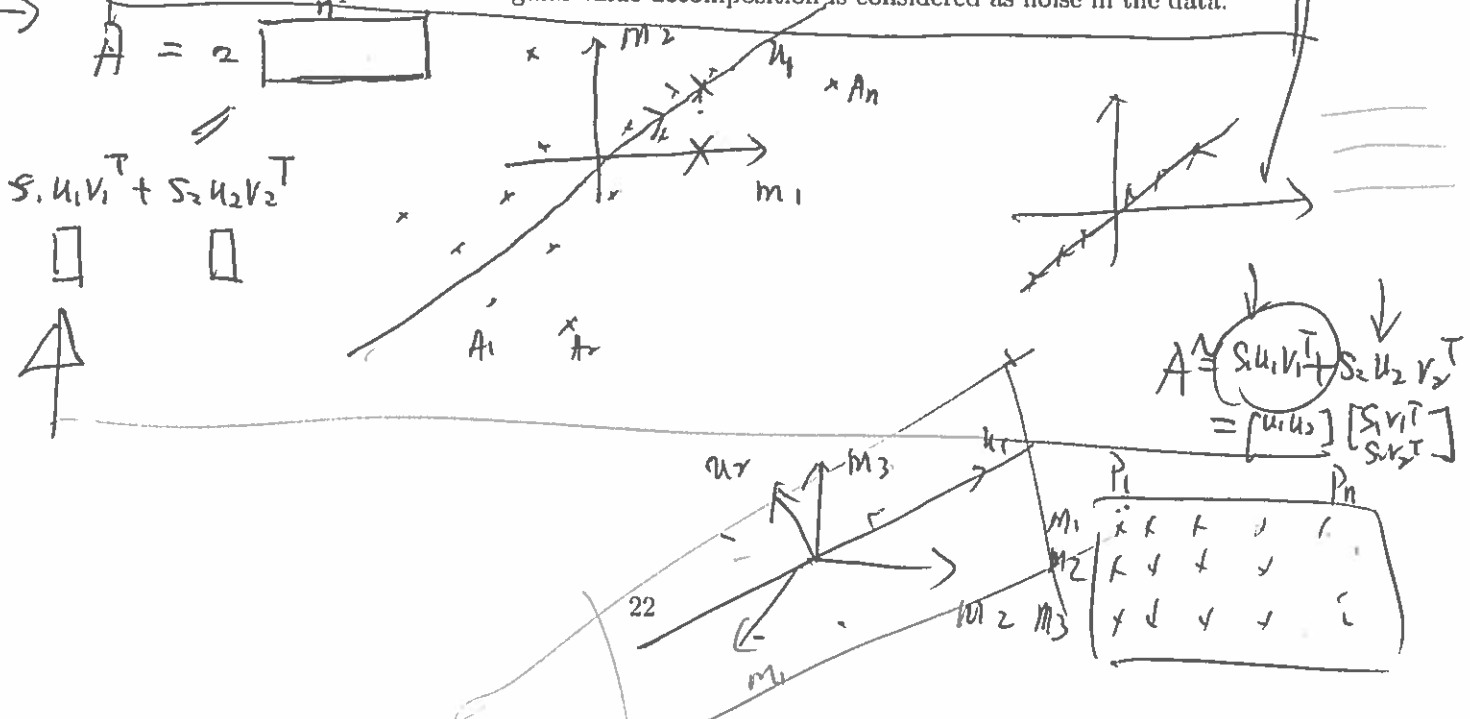
$$= \text{tr} AA^T / (n - 1) = T$$

$$T = \text{tr} A^T A / (n - 1) = \text{tr} S / (n - 1) = \left( \sum_{i=1}^r s_i^2 \right) / (n - 1)$$

where  $s_1 \geq \dots \geq s_r > 0$  are the singular values of  $A$ .

$$A = (s_1 u_1 v_1^T + \dots + s_k u_k v_k^T)$$

- The first  $k$  singular vectors capture more information than other vectors;  $u_j$  is referred to as the  $j$ th principal component of the data that accounts for the fraction  $s_j^2 / T$  of the variance.
- The effective rank  $k$  of  $A$  or  $S$  is the number of singular values larger than certain threshold so that the other part in the singular value decomposition is considered as noise in the data.



$$A = U \Sigma V^T$$

$$U^T A V = \Sigma$$

- Note that the line is different from finding the best fit  $y = ax + b$ . In that case, we want to find best  $a, b$  such that  $ax_i + b = y_i$  for  $i = 1, \dots, n$  without centering the data. We consider  $\tilde{A}(a, b)^T = (y_1, \dots, y_n)^T$  and find the least square solution:

$$\tilde{A}^T \tilde{A}(a, b)^T = \tilde{A}^T (y_1, \dots, y_n)^T.$$

This is known as standard least square.

- In our case, we consider the centered data, and

$$\|A^T\|_F^2 = \|A^T u_1\|_F^2 + \dots + \|A^T u_m\|_F^2$$

$$(\sum a_{ij}^2)$$

so that

the sum of squared distances from the data points to  $u_1, \dots, u_k$  is a minimum.

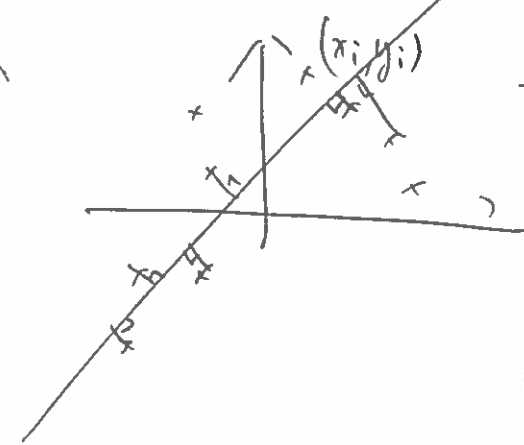
$$\left\| \begin{bmatrix} A_1^T \\ A_2^T \\ \vdots \\ A_n^T \end{bmatrix} \begin{bmatrix} u_1 \\ u_2 \end{bmatrix} \right\|_F$$

$$= \|A^T u_1\|_F^2 + \|\bar{A} u_2\|_F^2$$

There are interesting discussion of the Hilbert matrix

$$H = [a_{ij}] = [1/(i + j - 1)]$$

and the zero-one matrix representing the picture of square, triangle, circle, etc. See pp. 78-79.



$$y = ax + b$$

$$\begin{aligned} ax_1 + b &= y_1 \\ ax_2 + b &= y_2 \\ &\vdots \\ ax_n + b &= y_n \end{aligned}$$

$$\begin{bmatrix} x_1 & 1 \\ \vdots & \vdots \\ x_n & 1 \end{bmatrix} \begin{bmatrix} a \\ b \end{bmatrix} = \begin{bmatrix} y_1 \\ \vdots \\ y_n \end{bmatrix}$$

$$\hat{A} \begin{bmatrix} a \\ b \end{bmatrix} = \begin{bmatrix} y_1 \\ \vdots \\ y_n \end{bmatrix}$$

$$\hat{A}^T \hat{A} \begin{bmatrix} a \\ b \end{bmatrix} = \hat{A}^T \begin{bmatrix} y_1 \\ \vdots \\ y_n \end{bmatrix}$$

$$= \begin{bmatrix} x_1 & \dots & x_n \\ 1 & \dots & 1 \end{bmatrix} \begin{bmatrix} y_1 \\ \vdots \\ y_n \end{bmatrix}$$

Linear regression

$$H_1 = [1]$$

$$H_2 = \begin{bmatrix} 1 & 1/2 \\ 1/2 & 1/3 \end{bmatrix}$$

$$H_3 = \begin{bmatrix} 1 & 1/2 & 1/3 \\ 1/2 & 1/3 & 1/4 \\ 1/3 & 1/4 & 1/5 \end{bmatrix} = \sum_{i=1}^3 s_i u_i u_i^T$$

$$H_n = \sum s_i u_i u_i^T$$

If  $A = A^*$  is psd. then  $A = \lambda_1 u_1 u_1^* + \dots + \lambda_r u_r u_r^*$ .

$\lambda_1, \dots, \lambda_r$  are the positive eigenvalues.

$$U^* A U = \begin{bmatrix} \lambda_1 & & 0 \\ & \ddots & \\ 0 & & \lambda_r \\ \hline 0 & & 0 \end{bmatrix} = \lambda_1 u_1 u_1^* + \dots + \lambda_r u_r u_r^*$$

If  $A = A^*$  with eigenvalues  $\lambda_1 \geq \lambda_2 \geq \dots \geq \lambda_k > 0$   
&  $\lambda_n \leq \lambda_{n-1} \leq \dots \leq \lambda_{n-p+1} < 0$

$$\lambda_{k+1} = \dots = \lambda_{n-p} = 0$$

and  $A = \lambda_1 u_1 u_1^* + \dots + \lambda_k u_k u_k^* + \lambda_{n-p+1} u_{n-p+1} u_{n-p+1}^* + \dots + \lambda_n u_n u_n^*$

Then the SVD is:

$$\left\{ \begin{array}{l} u_1, \dots, u_k, u_{n-p+1}, \dots, u_n \\ \uparrow \\ u_1, \dots, u_k, v_{n-p+1}, \dots, v_n \end{array} \right\}$$

$$A = \lambda_1 u_1 u_1^* + \dots + \lambda_k u_k u_k^* + (\lambda_{n-p+1}) u_{n-p+1} (-u_{n-p+1})^* + \dots + (\lambda_n) u_n (-u_n)^*$$

$$\begin{array}{l} A^* A \\ \text{has e.v.} \\ |\lambda_1|^2, \dots, |\lambda_n|^2 \end{array}$$